

A SAS® Macro to Compute Added Predictive Ability of New Markers Predicting a Dichotomous Outcome

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ABSTRACT

Risk prediction is an important field in applied statistics. For example, in clinical research, predicting the development of adverse medical conditions is the object of many studies. However, this concept is multi-disciplinary. Many published models exist to predict dichotomous outcomes, ranging widely from a model to predict winners in the NCAA basketball tournament to a recently published risk model to predict the likelihood of a bleeding complication after percutaneous coronary intervention. It is important to note, however, these models are dynamic, as changes in population and technology lead to discovery of new and better methods to predict outcomes. An important consideration is when to add new markers to existing models. While a significant p-value is an important condition, it does not necessarily imply an improvement in model performance. Traditionally, receiver operating characteristic (ROC) curves and its corresponding area under the curve (AUC) are used to compare models, with a statistical test due to DeLong et al. for two correlated AUCs. However, the clinical relevance of this metric has been questioned by researchers^{1,2,3}. To address this issue, Pencina and D'Agostino¹ have proposed two statistics to evaluate the significance of novel markers. The Integrated Discrimination Improvement (IDI) measures the new model's improvement in average sensitivity without sacrificing average specificity. The Net Reclassification Improvement (NRI) measures the correctness of reclassification of subjects based on their predicted probabilities of events using the new model with the option of imposing meaningful risk categories.

Currently, only the AUC can be obtained directly from SAS (proc logistic). In contrast, the AUC as well as the NRI and IDI, can both be outputted in R with the ImproveProb function under the Hmisc library ([Click Here](#)). In this paper we present a single SAS macro which allows calculation of all three measures with several user-defined specifications.

INTRODUCTION

This paper is going to cover situations where we are comparing 2 models predicting a yes/no outcome without a time-to-event element. This can be done using a variety of different models, including; logistic, log-binomial, or Modified Poisson^{6,12} approach for relative risk estimation. For example, consider comparing 2 nested logistic regression models to predict a dichotomous outcome with the main difference being model 2 includes a variable for a new marker as a predictor:

$$\text{Model1} : \text{logit}(y) = \alpha + \beta_1 x_1 + \dots + \beta_n x_n$$

$$\text{Model2} : \text{logit}(y) = \alpha + \beta_1 x_1 + \dots + \beta_n x_n + \beta_{n+1} x_{n+1}$$

The output of each model will be a predicted probability of experiencing an event, and each individual in the dataset will have 2 probabilities, one from each model. Our goal is to evaluate how much the new variables to model2 add to prediction compared to model1.

AUC

Difference in the area under the receiver operating characteristic curve (AUC) is a common method to compare two models. AUC is a measure of model discrimination (i.e. how well the model separates subjects who did and did not experience an event). It essentially depicts a tradeoff between the benefit of a model (true positive or sensitivity) vs. its costs (false positive or 1-specificity). AUC is computed by comparing all possible pairs in a dataset between individuals experiencing and not experiencing an event. For example, a dataset with 100 events and 1000 non-events would have $100 \times 1000 = 100,000$ pairs. In each pair the predicted probability is compared. If the individual experiencing an event has a higher predicted probability, that pair would be labeled 'concordant' and assigned a value of 1. Conversely, if the individual experiencing an event has a lower probability the pair would be labeled 'discordant' and assigned a value of 0. AUC would be an average of the 1s and 0s (and 0.5s for identical values).

AUC ranges from 0.5 (no discrimination) to 1 (perfect discrimination); however, values close to 1 are very rarely seen in contemporary data^{3,7}. The value of baseline AUC is important, but in our context the focus is on the contribution of a new marker. To address this scenario, DeLong⁴ presented a method for comparing 2 correlated AUCs. Prior to SAS version 9.2 you could run the %roc macro located on the SAS website⁸. However, if you are running a logistic model in 9.2 this comparison can be done with the two new statements that were added to proc logistic (roc and roccontrast).

Sample code for SAS 9.2 Proc Logistic:

```
proc logistic data=auc;
model eve=x1 x2 x3 x4 x5;
roc 'first' x1 x2 x3 x4;
roc 'second' x1 x2 x3 x4 x5;
roccontrast reference('first')/estimate;
run;
```

The above code will return a comparison of model 1 (with the first 4 covariates) with model 2 (containing the additional covariate). For a pre-version 9.2 example go to the SAS website [\[Click Here\]](#).

Many advantages exist for reporting the AUC. First, it is a well known statistic, commonly seen in any model predicting a dichotomous outcome. Second, it is the default output in most all statistical software including proc logistic in SAS. Third, there are recommend ranges for an excellent/good/poor value for AUC making it easier to evaluate the model. In their textbook, Hosmer and Lemeshow⁷ consider AUC values of 0.7 to 0.8 to show acceptable discrimination, values of 0.8 to 0.9 to indicate excellent discrimination, and values of ≥ 0.9 to show outstanding discrimination. However, they also note how unusual it is to observe AUC values over .9 and how it could cause issues in the model due to complete separation of data.

There are also disadvantages to the AUC. First, it is a rank based statistic. A pair of individuals with probabilities 0.5 and 0.51 would be treated the same as with probabilities 0.1 and 0.9. Second, it focuses only on discrimination which is not the only and may not be the most important in the context of risk prediction. Third, it is very difficult for a new marker to significantly change the value of AUC³. Fourth, the magnitude of improvement in AUC is not nearly as meaningful as the value of AUC itself. For this reason it is very difficult to use increase in the AUC to justify adding new markers to a model even though these markers may have a major impact at the individual level³.

The next two measures (IDI/NRI) attempt to quantify the added predictive ability of a new marker by contributing additional information beyond the AUC. IDI attempts to add to AUC by quantifying how far apart in probabilities the events and non-events are when adding a new marker. And NRI attempts to quantify how many individuals are correctly "reclassified" when adding a new marker.

IDI

IDI is a measure of the new model's improvement in average sensitivity (true positive rate) without sacrificing its average specificity (true negative rate). In comparing the models, IDI measures the increment in the predicted probabilities for the subset experiencing an event and the decrement for the subset not experiencing an event (see figure1). This adds an important element that AUC lacks. The AUC is a rank based statistic in which all that matters is which probability is higher or lower, whereas IDI gives a measure of how far apart on average they are.

$$\text{Absolute_IDI} = (\bar{p}_{\text{event_2}} - \bar{p}_{\text{event_1}}) + (\bar{p}_{\text{nonevent_1}} - \bar{p}_{\text{nonevent_2}}) \quad (1)$$

$$\text{Relative_IDI} = \frac{\bar{p}_{\text{event_2}} - \bar{p}_{\text{nonevent_2}}}{\bar{p}_{\text{event_1}} - \bar{p}_{\text{nonevent_1}}} - 1 \quad (2)$$

where :

$\bar{p}_{\text{event_2}}$ is the mean *predicted probability* of the subset experiencing an event using model2

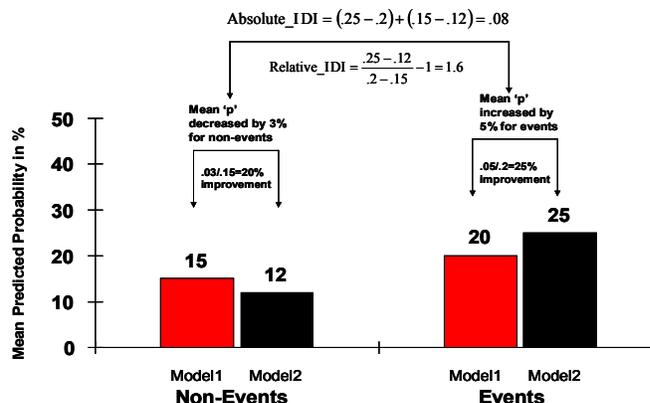


Figure 1: Graphical Depiction of IDI

One caveat is that absolute IDI depends on the event rate observed in a given data. If the event rate of interest is only 3% then the absolute IDI value will be seemingly small. This is where relative IDI is useful—it returns a percentage (no units). Pencina and D'Agostino provide an example where the absolute IDI is only 0.009 but

represents a 13% improvement on the relative scale. In Figure1 the relative IDI value of 1.6 indicates that the difference in predicted probabilities between events and non-events increased 1.6 times between model1 and model2.

NRI

NRI is based on a concept similar to IDI but its focus is on the upward and downward movement of predicted risks among those with and without events. Upward and downward movement can be defined based on simple increases and decreases in predicted risks using model2 instead model1, or applying categories and requiring movement between categories. NRI evaluates the 'net' number of individuals reclassified correctly using model2 over model1. This is done by calculating how many individuals experiencing an event increased in absolute risk or risk category (e.g., from medium to high) and how many individuals not experiencing an event decreased in absolute risk or risk category (e.g., medium to low).

$$NRI = \left[\left(\frac{\# \text{ Events moving up}}{\# \text{ of Events}} \right) - \left(\frac{\# \text{ Events moving down}}{\# \text{ of Events}} \right) \right] + \left[\left(\frac{\# \text{ Non - events moving down}}{\# \text{ of Non - events}} \right) - \left(\frac{\# \text{ Non - events moving up}}{\# \text{ of Non - events}} \right) \right] \quad (3)$$

The no category option can be seen as the categorical version with as many categories as there are estimated risks. In this definition any movement in probabilities counts as an up or down movement. This is the definition used by the ImproveProb function in R. Using this classification if an individual's probability using model 1 was 0.17 and using model 2 it became 0.18 this would be considered an upward movement. The category-based version involves a user defined classification. This requires more intimate knowledge of the data and population; for example one could define 3 groups: low (probability<10%), moderate (Probability 10-20%), and high (probability >20%). Here if an individual was a 'low' in model 1 but became a 'moderate' in model 2 it would be considered an upward movement.

There are advantages to both category-based and category-free NRI. With the category-free definition the user doesn't have to worry about determining meaningful categories, with the potential drawback of a loss of practical interpretation. The "user" defined classification can be easily seen in two nXn tables (see tables 1&2 below) if one imposes only a few groups, and can be expressed in more practical terms (eg. 10 people who had an event went from "medium" risk to "high" risk). Importantly, the developed macro has options to compute NRI using either or both of these definitions, a helpful addition to current statistical software packages.

NRI COMPUTATION EXAMPLE:

A dataset with 100 events and 200 non-events, using 3 user defined categories: <10% (low), 10-20% (moderate), >20% (high). By making 2 separate (event and non-event) crosstabs of model1 groups by model2 groups one can calculate NRI, by counting the numbers on the off-diagonal:

Table1: Crosstab for Events (n=100)

		Model 2			
		Low	Mod	High	
Model 1	Low	10	8	2	20 Events moving up 10 Events moving down 70 Events not moving Net of 10/100 (10%) of events getting reclassified correctly
	Mod	3	30	10	
	High	2	5	30	

Table2: Crosstab for Non-events (n=200)

		Model 2			
		Low	Mod	High	
Model 1	Low	50	5	0	15 Non-events moving up 35 Non-events moving down 150 Non-events not moving Net of 20/200 (10%) of non-events getting reclassified correctly
	Mod	20	40	10	
	High	5	10	60	

$$NRI = \left[\left(\frac{20}{100} \right) - \left(\frac{10}{100} \right) \right] + \left[\left(\frac{35}{200} \right) - \left(\frac{15}{200} \right) \right] = .2$$

It is important to remember that NRI is dependent on the groups chosen. For example, the results could be different had the groups been <15%, 15-30%, and >30%, by defining 4 groups instead of 3, or using category-free. Due to this issue, defining your NRI strategy before data analysis is essential..

ADDITIONAL MEASURES: CALIBRATION

So far this paper has discussed 3 important measures for evaluating the added predictive ability: increase in AUC, IDI, and NRI. However, one very important measure also needs to be addressed: calibration. Calibration quantifies how close the predicted probabilities match the actual experience. It is very important to when adding a new marker

to a model that there is at least no worsening in model calibration (hopefully we might see an improvement).

A traditional measure of calibration is the Hosmer-Lemeshow Chi-Square Test. This test breaks the data into “j” equal groups (frequently 10) based on ordered predicted probabilities and compares means of predicted predicted probabilities to observed rates. This test statistic can be compared to a Chi-square distribution with j-2 degrees of freedom, where a sufficiently large value will indicate a poorly calibrated model (eg, low p-values represent lack of fit)

$$HL_j = \sum_{j=1}^j \left[\frac{\left(\sum_{i=1}^{n_j} y_{ij} - \sum_{i=1}^{n_j} p_{ij} \right)^2}{\sum_{i=1}^{n_j} p_{ij} * (1 - \bar{p}_j)} \right] \sim \chi^2(j-2)$$

Where :

j is the number of groups (usually 10)

n_j is the number of subjects in the j 'th group

y_{ij} is the i 'th outcome of the j 'th group (0 or 1)

p_{ij} is the i 'th predicted probability of the j 'th group

\bar{p}_j is the mean predicted probability of the j 'th group

This test statistic has recently come under scrutiny by the original authors since it is driven by the number of groups and how they are defined. In a paper by Hosmer and Lemeshow they show that this test obtains different results from various statistical packages due to the algorithm used to select cutpoints. In the same paper they recommend other tests and the author refers the reader to this publication⁵. In [Appendix A](#) we provide a SAS macro to compute the HL test statistic with a default of 10 groups for any set of predicted probabilities and observed outcomes., This macro is then called in [Appendix B](#) for default output along with AUC/IDI/NRI. Alternatively, in SAS you can use proc logistic's "lackfit" option to compute the test statistic, however, the macro in appendix A is not limited to a logistic model. Importantly, there exists a SAS macro to compute the other statistics in the Hosmer paper and was presented at the 2001 SAS Global Forum (SUGI 26)⁸ using proc IML. Additionally the resid function in R computes one of the alternatives in the Hosmer & Lemeshow paper, unweighted sum of squares. ([Click Here](#)).

SAS MACRO TO OBTAIN OUTPUT

The entire code is in Appendix A&B and for more examples contact the corresponding author. In addition a worked macro call is provided in [Appendix C](#) using the Cars data from the sashelp library.

```
%macro add_predictive(data= , y= , p_old= , p_new=, nrpoints=%str(),  
hoslemgrp=10);
```

This macro only has 4 required user inputs. By entering just the data, outcome(y), initial probabilities, and final probabilities, the user will obtain output on AUC, IDI, NRI using the category-free definition, and the Hosmer-Lemeshow chi-square test (see appendix C for an example using the macro). Below are the optional inputs:

- 1) If the user wants defined NRI groups (i.e. <10%, 10-20%, and >20%) enter: **nrpoints=.1 .2**
- 2) If the user wants a different number of groups for the HL test (other than 10) enter: **hoslemgrp=#**

TIME TO EVENT ANALYSIS

It's important to note that this paper was set up under a model of predicting a dichotomous outcome without a time-to-event element. This was done for a few reasons. First, I intended this paper to serve as an introduction to a very important statistical concept and using this strategy allowed for ease of interpretation and explanation. Second, predicting a yes/no outcome is a very common situation in essentially all areas of research, making this macro applicable in many fields. Third, the new c-statistic options under version 9.2 in SAS with the ROC and ROCCONTRAST statements was important enough to warrant the restrictive case.

However, these or equivalent measures are available for time-to-event data. Interestingly, the initial paper on IDI/NRI was used analyzing the Framingham data ([Click Here](#)) which is a long term study; hence, time-to-event was used as the frame of the first discussion. Because of this, using predicted rates from survival analysis (eg. Mortality probability) could be inserted into the macro to get correct estimates of IDI and NRI. The difference comes in the calculation of the C-statistic and Hosmer-Lemeshow. Currently there is a comparable AUC measure for Survival data⁹ along with tests of correlated AUC measures¹⁰ and also papers to help compute survival AUC¹¹.

CONCLUSION

Predicting dichotomous events will continue to be an important aspect of research. Additionally, improving already useful models will be just as important. This paper presented 4 important measures in determining model performance and improvement by the addition of new markers. Pencina and D'Agostino¹ state, "the choice of the improvement metric should take into account the question to be answered". If the primary goal is to determine if an individual falls above or below a specific cut-off then NRI will probably be the preferred metric. On the other hand, if no cut-offs exist it may be best to look at category-free NRI, IDI and AUC. Also, the example from Pencina and D'Agostino shows a case where the addition of a new variable to a model results in a very small change in AUC but a marked improvement in IDI and NRI. The small change in AUC is explained more fully by Pepe et al.³ and may justify not solely using increase in AUC as the guideline for model improvement but rather a combination of metrics. It is also important to note that other factors will play a role in determining whether to add a new marker. For example, if the new marker is more difficult and expensive to obtain guidelines on what constitutes a significant improvement in model performance will likely be stricter. It is important to keep in mind all of these factors in assessing model performance.

REFERENCES

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CONTACT INFORMATION

Your comments are very encouraged please contact me at

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APPENDIX A: HOSMER LEMESHOW MACRO

```
%macro hoslem (data=, pred=, y=, ngro=10,print=T,out=hl);
/*---
Macro computes the Hosmer-Lemeshow Chi Square Statistic
for Calibration

Parameters (* = required)
-----
data*          input dataset
pred*          Predicted Probabilities
y*            Outcome 0/1 variable
ngro          # of groups for the calibration test (default 10)
print         Prints output (set to F to Suppress)
out           output dataset (default HL)

Author: Kevin Kennedy
---*/
%let print = %upcase(&print);
proc format;
value pval 0-.0001='<.0001';
run;

data first;set &data;where &y^=. and &pred^=.;run;
proc rank groups=&ngro out=ranks data=first;ranks phat_grp;var &pred;run;
proc sort data=ranks;by phat_grp;run;
proc sql;
create table ranks2 as select *, count(*) as num_dec label='Sample Size', sum(&pred) as sum_pred
label='Sum Probabilities', sum(&y) as sum_y label='Number of Events'
from ranks
group by phat_grp;
create table ranks3 as select distinct(phat_grp),num_dec,sum_pred,sum_y,
((sum_y-sum_pred)**2/(sum_pred*(1-sum_pred/num_dec))) as chi_part label 'Chi-Square Term'
from ranks2 ;
select sum(chi_part) into :chi_sq
from ranks3;
quit;

data &out;
chi_sq=&chi_sq; label chi_sq='Hosmer Lemeshow Chi Square';
df=&ngro-2; label df='Degree of Freedom';
p_value=1-cdf('chisquared',chi_sq,df);label p_value='P-Value';
format p_value pval.;
run;
%if &print=T %then %do;
title 'Hosmer Lemeshow Details';
proc print data=ranks3 noobs label;run;
Options formdlim='-';
title 'Hosmer Lemeshow Calibration Test';
proc print data=&out noobs label;run;
%end;
Options formdlim='';
%mend;
```

APPENDIX B: ADD_PREDICTIVE MACRO (MACRO CALLS %HOSLEM IN APPENDIX A)

```
%macro add_predictive(data=, y=, p_old=, p_new=, nripoints=%str(),hoslemgrp=10);

/*---
this macro attempts to quantify the added predictive ability of a covariate(s)
in logistic regression based off of the statistics in: M. J. Pencina ET AL. Statistics
in Medicine (2008 27(2):157-72). Statistics returned with be: C-statistics (for 9.2 users),
IDI (INTEGRATED DISCRIMINATION IMPROVEMENT), NRI (net reclassification index)
for both Category-Free and User defined groups, and Hosmer Lemeshow GOF test
with associated pvalues and z scores.

Parameters (* = required)
-----
data*           Specifies the SAS dataset

y*             Response variable (Outcome must be 0/1)

p_old*        Predicted Probability of an Event using Initial Model

p_new*        Predicted Probability of an Event using New Model

nripoints      Groups for User defined classification (Optional),
Example 3 groups: (<.06, .06-.2, >.2) then nripoints=.06 .2

hoslemgrp      # of groups for the Hosmer Lemeshow test (default 10)

Author: Kevin Kennedy and Michael Pencina
Date: May 26, 2010

---*/

options nonotes nodate nonumber;
ods select none;
%local start end;
%let start=%sysfunc(datetime());

proc format;
value pval 0-.0001='<.0001';
run;

/*****Step 1: C-Statistics*****/
/*****Running AUC Analysis*****/
%if %sysevalf(&sysver >= 9.2) %then %do;
%put *****Running AUC Analysis*****;
proc logistic data=&data descending;
model &y=&p_old &p_new;
roc 'first' &p_old;
roc 'second' &p_new;
rocontrast reference('first')/estimate e;
ods output ROCAssociation=rocass ROCContrastEstimate=rocdiff;
run;

proc sql noprint;
select estimate, StdErr, lowercl, uppercl, (ProbChiSq*100) as pval
into :rocdiff, :rocdiff_stderr, :rocdiff_low, :rocdiff_up, :rocp
from rocdiff
where find(contrast, 'second');
quit;

data _null_;
set rocass;
if ROCModel='first' then do;
call symputx('c_old',Area);
end;
if ROCModel='second' then do;
call symputx('c_new',Area);
end;
run;

data cstat;
cstat_old=&c_old; label cstat_old='Model1 AUC';
cstat_new=&c_new; label cstat_new='Model2 AUC';
cstat_diff=&rocdiff; label cstat_diff='Difference in AUC';
cstat_stderr=&rocdiff_stderr; label cstat_stderr='Standard Error of Difference in AUC';
cstat_low=&rocdiff_low; label cstat_low='Difference in AUC Lower 95% CI';
```

```

cstat_up=&rocdiff up; label cstat_up='Difference in ACU Upper 95% CI';
cstat_ci='(!trim(left(cstat low))!','!trim(left(cstat_up))!');
label cstat_ci='95% CI for Difference in AUC';
cstat_pval=&rocp/100; label cstat_pval='P-value for AUC Difference';
format cstat_pval pval.;
run;

%end;
%if %sysevalf(&sysver < 9.2) %then %do;
options notes ;
%put *****;
%put NOTE: You are running a Pre 9.2 version of SAS;
%put NOTE: Go to SAS website to get example of ROC Macro for AUC Comps;
%put NOTE: http://support.sas.com/kb/25/017.html;
%put *****;
%put;
options nonotes ;
%end;

/*****End step 1*****/
/*****Step 2: IDI*****/
%put *****Running IDI Analysis*****;
proc sql noprint;
create table idinri as select &y, &p_old, &p_new, (&p_new-&p_old) as pdiff
from &data
where &p_old^=. and &p_new^=.
order by &y;
quit;

proc sql noprint; /*define mean probabilities for old and new model and event and nonevent*/
select count(*), avg(&p_old), avg(&p_new), stderr(pdiff) into
:num_event, :p_event_old, :p_event_new, :eventstderr
from idinri
where &y=1 ;
select count(*), avg(&p_old), avg(&p_new), stderr(pdiff) into
:num_nonevent, :p_nonevent_old, :p_nonevent_new, :noneventstderr
from idinri
where &y=0;
quit;

data fin(drop=slope noadd slope add);
pen=&p_event new; label pen='Mean Probability for Events: Model2';
peo=&p_event_old; label peo='Mean Probability for Events: Model1';
pnen=&p_nonevent_new; label pnen='Mean Probability for NonEvents: Model2';
pneo=&p_nonevent_old; label pneo='Mean Probability for NonEvents: Model1';
idi=(&p_event_new-&p_nonevent_new)-(&p_event_old-&p_nonevent_old);
label idi='Integrated Discrimination Improvement';
idi_stderr=sqrt((&eventstderr**2)+(&noneventstderr**2));
label idi_stderr='IDI Standard Error';
idi_lowci=round(idi-1.96*idi_stderr, .0001);
idi_upci=round(idi+1.96*idi_stderr, .0001);
idi_ci='(!trim(left(idi_lowci))!','!trim(left(idi_upci))!');
label idi_ci='IDI 95% CI';
z_idi=abs(idi/(sqrt((&eventstderr**2)+(&noneventstderr**2))));
label z_idi='Z-value for IDI';
pvalue_idi=2*(1-PROBNORM(abs(z_idi))); label pvalue_idi='P-value for IDI';
change_event=&p_event_new-&p_event_old;
label change_event='Probability change for Events';
change_nonevent=&p_nonevent_new-&p_nonevent_old;
label change_nonevent='Probability change for Nonevents';
slope_noadd=&p_event_old-&p_nonevent_old;
slope_add=&p_event_new-&p_nonevent_new;
relative_idi=slope_add/slope_noadd-1; label relative_idi='Relative IDI';
format pvalue_idi pval.;
run;

/*****step 3 NRI analysis*****/
%put *****Running NRI Analysis*****;
data nri inf;
set idinri;
if &y=1 then do;
down_event=(pdiff<0); up_event=(pdiff>0); down_nonevent=0; up_nonevent=0;
end;
if &y=0 then do;
down_nonevent=(pdiff<0); up_nonevent=(pdiff>0); down_event=0; up_event=0;
end;

```

```

run;

proc sql;
select sum(up_nonevent), sum(down_nonevent), sum(up_event),sum(down_event)
into :num_nonevent_up_user, :num_nonevent_down_user, :num_event_up_user, :num_event_down_user
from nri inf
quit;
/* Category-Free Groups */
data nril;
group="Category-Free NRI";
p_up_event=&num_event_up_user/&num_event;
p_down_event=&num_event_down_user/&num_event;
p_up_nonevent=&num_nonevent_up_user/&num_nonevent;
p_down_nonevent=&num_nonevent_down_user/&num_nonevent;

nri=(p_up_event-p_down_event)-(p_up_nonevent-p_down_nonevent);
nri_stderr=sqrt(((&num_event_up_user+&num_event_down_user)/&num_event**2-(&num_event_up_user-
&num_event_down_user)**2/&num_event**3)+
((&num_nonevent_down_user+&num_nonevent_up_user)/&num_nonevent**2-
(&num_nonevent_down_user-&num_nonevent_up_user)**2/&num_nonevent**3));
low_nrici=round(nri-1.96*nri_stderr,.0001);
up_nrici=round(nri+1.96*nri_stderr,.0001);
nri_ci='(!trim(left(low_nrici))!','!trim(left(up_nrici))!')';
z_nri=nri/sqrt(((p_up_event+p_down_event)/&num_event)
+((p_up_nonevent+p_down_nonevent)/&num_nonevent));
pvalue_nri=2*(1-PROBNORM(abs(z_nri)));
event_correct_reclass=p_up_event-p_down_event;
nonevent_correct_reclass=p_down_nonevent-p_up_nonevent;
z_event=event_correct_reclass/sqrt((p_up_event+p_down_event)/&num_event);
pvalue_event=2*(1-probnorm(abs(z_event)));
z_nonevent=nonevent_correct_reclass/sqrt((p_up_nonevent+p_down_nonevent)/&num_nonevent);
pvalue_nonevent=2*(1-probnorm(abs(z_nonevent)));
format pvalue_nri pvalue_event pvalue_nonevent pval. event_correct_reclass
nonevent_correct_reclass percent.;
label nri='Net Reclassification Improvement'
nri_stderr='NRI Standard Error'
low_nrici='NRI lower 95% CI'
up_nrici='NRI upper 95% CI'
nri_ci='NRI 95% CI'
z_nri='Z-Value for NRI'
pvalue_nri='NRI P-Value'
pvalue_event='Event P-Value'
pvalue_nonevent='Non-Event P-Value'
event_correct_reclass='% of Events correctly reclassified'
nonevent_correct_reclass='% of Nonevents correctly reclassified';

run;

/*User Defined NRI*/
%if &nripoints^=%str() %then %do;
/*words macro*/
%macro words(list,delim=%str( ));
%local count;
%let count=0;
%do %while(%qscan(%bquote(&list),&count+1,%str(&delim)) ne %str());
%let count=%eval(&count+1);
%end;
&count
%mend words;
%let numgroups=%eval(%words(&nripoints)+1); /*figure out how many ordinal groups*/
proc format ;
value group
1 = "0 to %scan(&nripoints,1,%str( ))"
%do i=2 %to %eval(&numgroups-1);
%let j=%eval(&i-1);
&i="%scan(&nripoints,&j,%str( )) to %scan(&nripoints,&i,%str( ))"
%end;
%let j=%eval(&numgroups-1);
&numgroups="%scan(&nripoints,&j,%str( )) to 1";
run;

data idinri;
set idinri;
/*define first ordinal group for pre and post*/
if 0<=&p_old<=%scan(&nripoints,1,%str( )) then group_pre=1;
if 0<=&p_new<=%scan(&nripoints,1,%str( )) then group_post=1;

%let i=1;
%do %until(&i>%eval(&numgroups-1));

```

```

        if %scan(&nripoints,&i,%str( ))<&p_old then do;
            group_pre=&i+1;
        end;
        if %scan(&nripoints,&i,%str( ))<&p_new then do;
            group_post=&i+1;
        end;
        %let i=%eval(&i+1);
    %end;
    if &y=0 then do;
        up_nonevent=(group_post>group_pre);
        down_nonevent=(group_post<group_pre);
        down_event=0; up_event=0;
    end;
    if &y=1 then do;
        up_event=(group_post>group_pre);
        down_event=(group_post<group_pre);
        down_nonevent=0; up_nonevent=0;
    end;
    format group_pre group_post group.;
run;

proc sql;
select sum(up nonevent), sum(down nonevent), sum(up event),sum(down event),avg(&y)
into :num_nonevent_up_user, :num_nonevent_down_user, :num_event_up_user,
:num_event_down_user, :eventrate
from idinri
quit;

data nri2;
group='User Category NRI';
p_up_event=&num_event_up_user/&num_event;
p_down_event=&num_event_down_user/&num_event;
p_up_nonevent=&num_nonevent_up_user/&num_nonevent;
p_down_nonevent=&num_nonevent_down_user/&num_nonevent;

nri=(p_up_event-p_down_event)-(p_up_nonevent-p_down_nonevent);
nri_stderr=sqrt(((&num_event_up_user+&num_event_down_user)/&num_event**2-
(&num_event_up_user-&num_event_down_user)**2/&num_event**3)+
((&num_nonevent_down_user+&num_nonevent_up_user)/&num_nonevent**2-
(&num_nonevent_down_user-&num_nonevent_up_user)**2/&num_nonevent**3));
low_nrici=round(nri-1.96*nri_stderr,.0001);
up_nrici=round(nri+1.96*nri_stderr,.0001);
nri_ci='(!!trim(left(low_nrici))!!','!!trim(left(up_nrici))!!)';
z_nri=nri/sqrt(((p_up_event+p_down_event)/&num_event)
+(p_up_nonevent+p_down_nonevent)/&num_nonevent));
pvalue_nri=2*(1-PROBNORM(abs(z_nri)));
event_correct_reclass=p_up_event-p_down_event;
nonevent_correct_reclass=p_down_nonevent-p_up_nonevent;
z_event=event_correct_reclass/sqrt((p_up_event+p_down_event)/&num_event);
pvalue_event=2*(1-probnorm(abs(z_event)));
z_nonevent=nonevent_correct_reclass/sqrt((p_up_nonevent+p_down_nonevent)/&num_nonevent);
pvalue_nonevent=2*(1-probnorm(abs(z_nonevent)));
format pvalue_nri pval.;
run;

data nril;
set nril nri2;
run;
%end;

/*****
/*step 4 gof */
*****/
%hoslem(data=idinri,pred=&p_old,y=&y,ngro=&hoslemgrp,out=m1,print=F);
%hoslem(data=idinri,pred=&p_new,y=&y,ngro=&hoslemgrp,out=m2,print=F);
data hoslem(drop=cnt);
retain model;
set m1 m2;
cnt+1;
if cnt=1 then model='Model1';
else model='Model2';
run;
ods select all;
/*output for cstat*/
%if %sysevalf(&sysver >= 9.2) %then %do;
proc print data=cstat label noobs;
title1 "Evaluating added predictive ability of model2";
title2 'AUC Analysis';run;

```

```

%END;
/*output for IDI*/
proc print data=fin label noobs;
title1 "Evaluating added predictive ability of model2";
title2 'IDI Analysis';
var idi idi_stderr z_idi pvalue_idi idi_ci
pen peo pnen pneo change_event change_nonevent relative_idi;
run;
/*output for NRI*/
proc print data=nri1 label noobs;
title1 "Evaluating added predictive ability of model2";
title2 'NRI Analysis';
var group nri nri_stderr z_nri pvalue_nri nri_ci event_correct_reclass pvalue_event
nonevent_correct_reclass pvalue_nonevent;
run;

%if &nripoints^=%str() %then %do;
  proc freq data=idinri;
  where &y=0;
  title 'NRI Table for Non-Events';
  tables group_pre*group_post/nopercent nocol;
  run;
  proc freq data=idinri;
  where &y=1;
  title 'NRI Table for Events';
  tables group_pre*group_post/nopercent nocol;
  run;
%end;
/*print HL gof*/
proc print data=hoslem noobs label;
title "Hosmer Lemeshow Test with %sysevalf(&hoslemgrp-2) df";
run;
proc datasets library=work nolist;
delete fin idinri nri1 nri2 nri_inf stderr;
quit;
options notes;
%put NOTE: Macro %nrstr(%%)add_predictive completed.;
%let end=%sysfunc(datetime());
%let runtime=%sysfunc(round(%sysevalf(&end-&start)));
%put NOTE: Macro Real Run Time=&runtime seconds;
title;
%mend;

```

APPENDIX C: EXAMPLE

```

data cars;
set sashelp.cars;
msrp_40k=(msrp>40000); /*indicator for car being over $40,000*/
cnt+1;
run;

/*Test to see if Country of Origin adds to the prediction of
msrp_40k with engine size, weight, and
MPG_Highway in model

Using NRI groups (<10%, 10-30%, >30%)*

/*initial model*/
proc logistic data=cars descending;
model msrp_40k=engine size weight mpg_highway;
output out=m1 pred=p1;
run;

/*new model*/
proc logistic data=cars descending;
class origin;
model msrp_40k=engine size weight mpg_highway origin;
output out=m2 pred=p2;
run;

proc sql;
create table cars2 as select *
from m1 as a left join m2 as b on a.cnt=b.cnt;
quit;

%add_predictive(data=cars2,y=msrp_40k,p_old=p1,p_new=p2, nripoints=.1 .3);

/*****
/*****Output*****/
/*****/

```

1) AUC

Model1 AUC	Model2 AUC	Difference in AUC	Standard Error of Difference in AUC	Difference in AUC Lower 95% CI	Difference in ACU Upper 95% CI	95% CI for Difference in AUC	P-value for AUC Difference
0.81748	0.93173	0.1142	0.0211	0.0729	0.1556	(0.0729,0.1556)	<.0001

2) IDI

Integrated Discrimination Improvement	IDI Standard Error	Z-value for IDI	P-value for IDI	IDI 95% CI	Mean Probability for Events: Model2	Mean Probability for Events: Model1
0.25235	0.028064	8.99191	<.0001	(0.1973,0.3074)	0.63917	0.44696

Mean Probability for NonEvents: Model2	Mean Probability for NonEvents: Model1	Probability change for Events	Probability change for Nonevents	Relative IDI
0.1129	0.17304	0.19221	-0.060139	0.92126

3) NRI

group	Net Reclassification Improvement	NRI Standard Error	Z-Value for NRI	NRI PValue	NRI 95% CI	% of Events correctly reclassified	Event PValue	% of Nonevents correctly reclassified	Non-Event PValue
Category-Free NRI	0.99832	0.10154	8.79944	<.0001	(0.7993,1.1973)	37%	0.0002	63%	<.0001
User Category NRI	0.44148	0.06793	6.13591	<.0001	(0.3083,0.5746)	20%	0.0012	25%	<.0001

3a) 3x3 table for User NRI—Non-Events

Table of group_pre by group_post				
group_pre	group_post			
Frequency Row Pct	0 to .1	.1 to .3	.3 to 1	Total
0 to .1	132 85.16	23 14.84	0 0.00	155
.1 to .3	85 74.56	11 9.65	18 15.79	114
.3 to 1	19 33.33	17 29.82	21 36.84	57
Total	236	51	39	326

3b) 3x3 table for User NRI—Events

Table of group_pre by group_post				
group_pre	group_post			
Frequency Row Pct	0 to .1	.1 to .3	.3 to 1	Total
0 to .1	1 14.29	5 71.43	1 14.29	7
.1 to .3	4 14.81	0 0.00	23 85.19	27
.3 to 1	3 4.41	2 2.94	63 92.65	68
Total	8	7	87	102

4) Hosmer-Lemeshow deciles test

model	Hosmer Lemeshow Chi Square	Degrees of Freedom	P-Value
Model1	16.7394	8	0.03294
Model2	7.4034	8	0.49379